

Optimal Dynamic Majority-Quorum Rules: Theory and Evidence

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Collective decisions trade off accuracy against time

◀ act too early

too little evidence

act too late ▶

too much delay

DAO governance

enact an opposed upgrade

prolong a vulnerability window

Crowd annotation

mislabel the item

burn budget, slow the pipeline

Election audits

certify the wrong winner

delay the transfer of power

The value of a decision depends on *what* is chosen and *when*.

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- In many modern settings the path of votes is observable, not just the final tally
- Rules could **adapt to the evolving informational state** — instead of fixed voting windows, participation thresholds, or predetermined sample sizes

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We study **optimal collective decision rules**: rules that identify the majority preference by acting on the observed vote stream.

- Three structural features define the problem

finite, exhaustible pool

vs. Wald-style tests: inexhaustible signal source

exogenous arrivals

vs. dynamic voting design: planner solicits votes at will

hard deadline

vs. finite-population audits: no binding horizon

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Related literature

- **Statistical social choice**: votes as noisy signals of a latent quantity
→ we keep the inference view, but the **rule is the choice variable**
Condorcet 1785; Young '88; Nitzan–Paroush '84; Austen-Smith–Banks '96;
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→ here: the sample is **given**, not built.
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Environment

- Society of N voters (N odd), each with a fixed binary preference: $+$ or $-$
- Unknown number of supporters $M^+ \in \{0, \dots, N\}$, $M^- = N - M^+$
- Benevolent planner: accept iff $M^+ > N/2$
- Uniform prior $M^+ \sim \text{Unif}\{0, \dots, N\}$

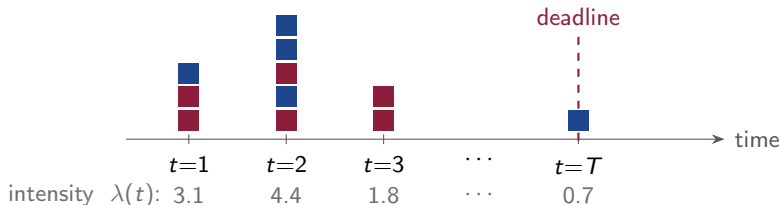
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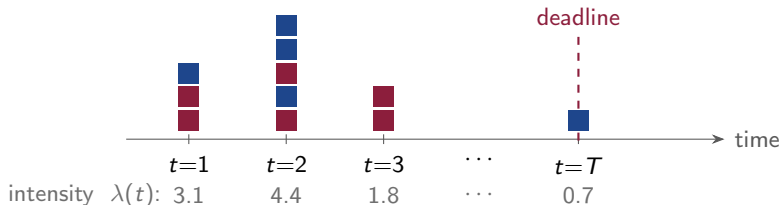
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Sequential vote arrival



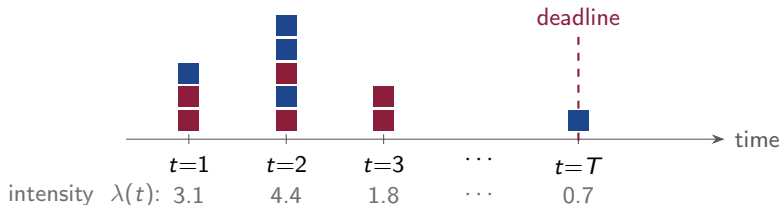
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- Cumulative counts s_t^+, s_t^- , turnout $s_t = s_t^+ + s_t^-$, lead $d_t = |s_t^+ - s_t^-|$
- Batch sizes: $A_t \sim \text{Poisson}(\lambda(t))$ independent across t ; $\Lambda(t) = \sum_{\tau \leq t} \lambda(\tau)$
- Each voter votes at most once $\Rightarrow$ censoring $a_t = \min\{A_t, N - s_{t-1}\}$
- Sampling **without replacement**, $a^+ | a, M^+ \sim \text{Hypergeom}(N - s, M^+ - s^+, a)$

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Learning

- Voters exchangeable $\Rightarrow (s_t^+, s_t^-)$ is a **sufficient statistic**

Lemma (Predictive distribution)

For a batch of size $a \leq N - s$,

$$P(a^+ \mid a, s^+, s^-) = \frac{\binom{a^+ + s^+}{s^+} \binom{a - a^+ + s^-}{s^-}}{\binom{a + s + 1}{a}}, \quad a^+ \in \{0, \dots, a\}.$$

- Beta-Binomial($a, s^+ + 1, s^- + 1$): independent of N , which enters only via censoring and the majority cutoff
- Posterior belief $\mu(s^+, s^-) = P(M^+ > \frac{N}{2} \mid s^+, s^-) = \sum_{m^+ > N/2} \frac{\binom{m^+}{s^+} \binom{N - m^+}{s^-}}{\binom{N + 1}{s + 1}}$
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Planner's problem

At each t , after observing (s_t^+, s_t^-) : **stop** or **continue** at flow cost $\kappa > 0$.

- Stopping payoff (implement the leading alternative):

$$U(s^+, s^-) = R \cdot \max\{\mu(s^+, s^-), 1 - \mu(s^+, s^-)\}, \quad R > 0$$

- Bellman equation, terminal condition $V_T = U$:

$$V_t(s^+, s^-) = \max\left\{ U(s^+, s^-), \underbrace{-\kappa + \mathbb{E}[V_{t+1}(\tilde{s}_{t+1}^+, \tilde{s}_{t+1}^-) \mid s^+, s^-]}_{C_t(s^+, s^-)} \right\}$$

expectation over censored-Poisson batch size and Beta–Binomial composition

- Dynamic voting rule $\phi : (s^+, s^-, t) \mapsto \{\text{stop}, \text{continue}\}$; ϕ^* **optimal**

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The optimal rule is a dynamic majority-quorum rule

Theorem (Threshold structure)

The optimal voting rule ϕ^ is a dynamic majority-quorum rule: for all $(s_t^+, s_t^-) \in \mathcal{S}$ and $t \in \{1, \dots, T\}$ there exists a unique threshold $d_t^*(s_t) \geq 0$ such that ϕ^* prescribes continuation if and only if $|s_t^+ - s_t^-| < d_t^*(s_t)$. When $s_t = N$, $d_t^*(N) = 0$.*

- The threshold depends on the lead, turnout, *and* time; a given lead may be decisive at high turnout but not at low
- **Proof idea:** The proof establishes that a larger current lead makes future beliefs stochastically more favorable. Consequently, the net gain from stopping is monotone in the current lead. [▶ details](#)

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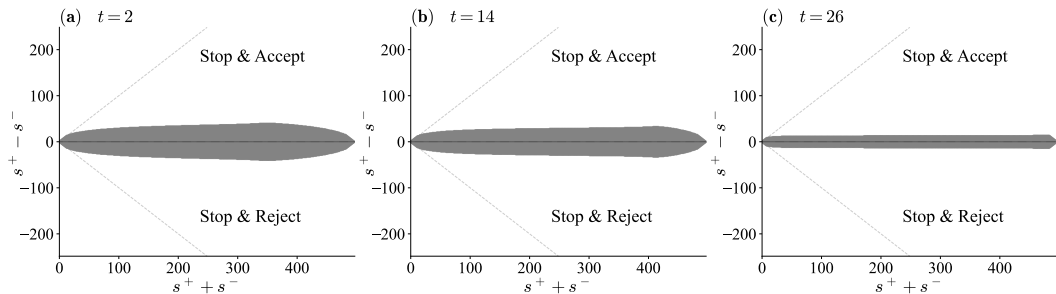
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The exact boundary: a collapsing lens



Exact rule ϕ^* in lead-turnout space at several dates; shaded: waiting is optimal. $N = 497$, $T = 28$, $\kappa = 0.001$, $R = 100$ ($\rho = 10^5$), uniform arrivals with $\Lambda(T) = 159$.

- **Two forces shape the lens:** statistical precision (small leads informative only at sufficient turnout $\Rightarrow$ widening) vs. finite-pool depletion (surviving leads become decisive as $s_t \rightarrow N \Rightarrow$ collapse), while the whole boundary sinks as $t \rightarrow T$

Comparative statics

Proposition

Under ϕ^* , the threshold $d_t^*(s; \rho, \lambda)$ with $\rho \equiv R/\kappa$ satisfies:

- i) $d_t^*(s; \rho', \lambda) \leq d_t^*(s; \rho, \lambda)$ for all $\rho' \leq \rho$, t , s ;
- ii) $d_t^*(s; \rho, \lambda') \leq d_t^*(s; \rho, \lambda)$ whenever $\lambda'(t) \leq \lambda(t)$ for all $t \in \{1, \dots, T\}$;
- iii) $d_{t+1}^*(s; \rho, \lambda) \leq d_t^*(s; \rho, \lambda)$ for all $t < T$, s , whenever $\lambda(\tau) \geq \lambda(\tau + 1)$ for all $\tau \in \{t + 1, \dots, T - 1\}$.

- One common force: the **option value of continued sampling**
- The conditions are tight: cumulative dominance does *not* suffice in (ii), and back-loaded arrivals can reverse monotonicity in (iii) ▶ counterexamples

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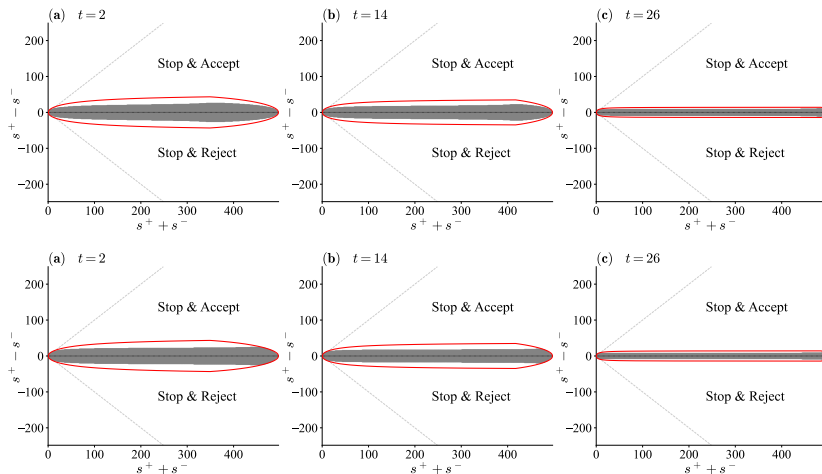
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Illustration: how the boundary moves with ρ and λ



Red curve: **benchmark** parameterization.

Lower stakes (top) and thinner arrivals (bottom) each **compress** the continuation lens, as in parts (i)–(ii).

From backward induction to a mechanism

- No closed form for $d_t^*(s)$: the value of waiting is defined recursively through later stopping decisions, on a two-dimensional state, over a finite horizon, with a transition law that shifts as the pool depletes
- Backward induction is exact but is too opaque to be used as a **governance mechanism**
- Strategy: derive an approximate closed-form rule

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A closed-form rule with a constant threshold

Censored forecast of terminal turnout: $\hat{s}_T := \min\{s_t + (\Lambda(T) - \Lambda(t)), N\}$.

Definition (Adaptive rule)

Let $\hat{\phi}$ be the majority-quorum rule stopping at (d_t, s_t, t) iff $t = T$, or $s_t = N$, or $t < T$ and

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- LHS: squared **signal-to-noise ratio** — strength of the sample lead over the **reducible** sampling variance: current noise $1/s_t$ *net of* the noise $1/\hat{s}_T$ that survives to the deadline
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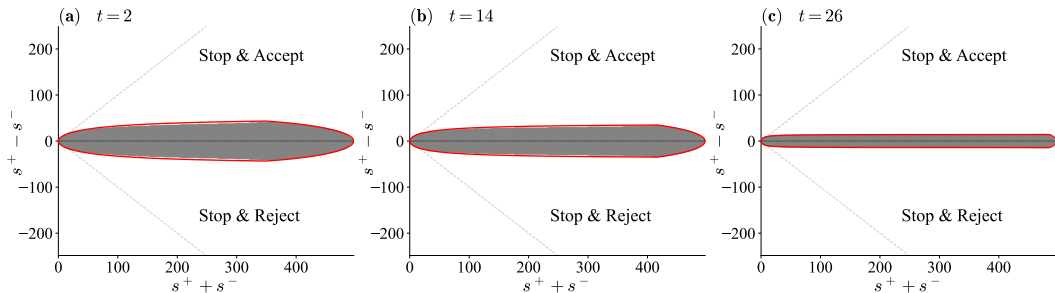
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The closed form tracks the exact boundary



Exact rule ϕ^* in lead-turnout space at several dates and $\hat{\phi}$ (red); $N = 497$, $T = 28$, $\kappa = 0.001$, $R = 100$ ($\rho = 10^5$), uniform arrivals with $\Lambda(T) = 159$.

The adaptive rule replicates the optimal policy

Theorem (Approximation)

Fix T . For every $\varepsilon \in (0, \frac{1}{2})$ there exist $N_\varepsilon, \rho_\varepsilon$ such that, for all $N \geq N_\varepsilon$ and $\rho \geq \rho_\varepsilon$ with $\log \rho \leq \varepsilon N^{1/3}$, ϕ^* and $\hat{\phi}$ prescribe the same action at every state $(d_t, s_t, t) \in \mathcal{S}_\varepsilon$ at which

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$\hat{\phi}$ cannot be represented as a dated schedule; the factor $N/(N - s_t)$ cancels in ratio form; carrying it through and evaluating the threshold on the mean path $s_t \rightarrow \Lambda(t)$ fixes it ex ante:

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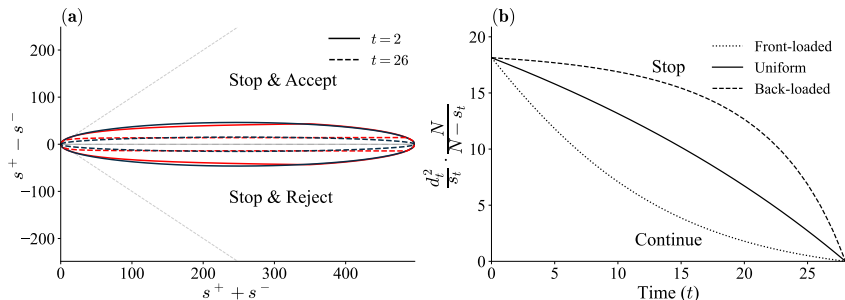
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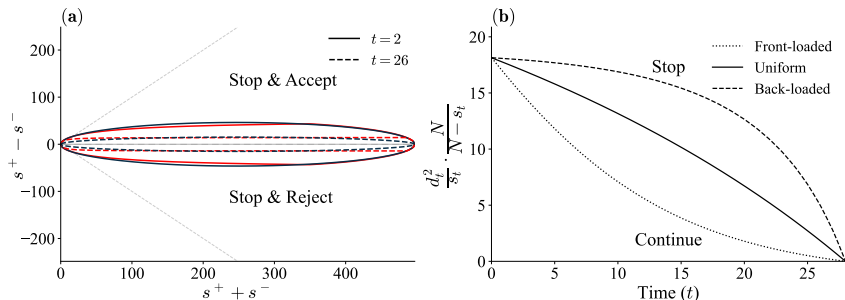
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The two rules, and the shape of the collapse



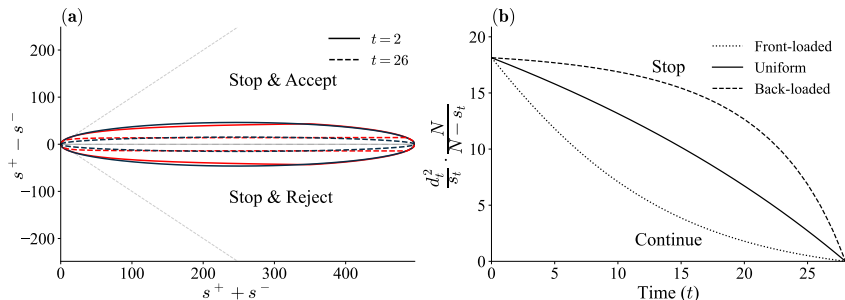
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The rule as a sequential significance test

$\rho = R/\kappa$ is hard to reason about in practice; restate the same rule via an error bound:

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- Among the largest DAOs by market cap; all protocol changes and treasury spending via **binding on-chain referenda**, 28-day decision periods
- Full vote path observed at twice-daily resolution $\rightarrow$ maps directly to (s_t^+, s_t^-)
- Incumbent rule is itself dynamic: the Approval-Support Mechanism (ASM)
- $\Rightarrow$ strategic timing is possible and likely
- Data: 1,824 referenda (06/2023–10/2025) $\rightarrow$ 952 cleaned $\rightarrow$ 933 in the stopping evaluation
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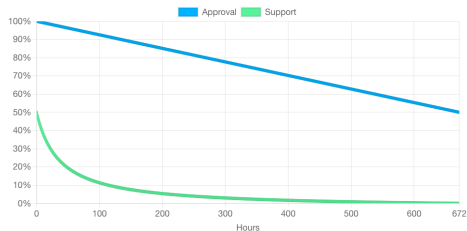
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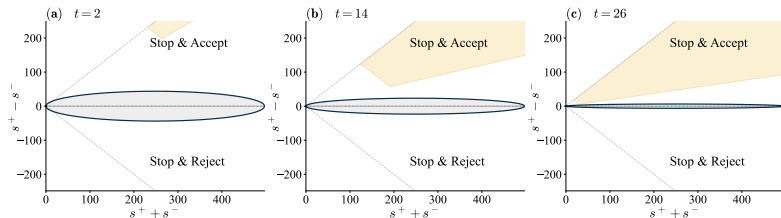
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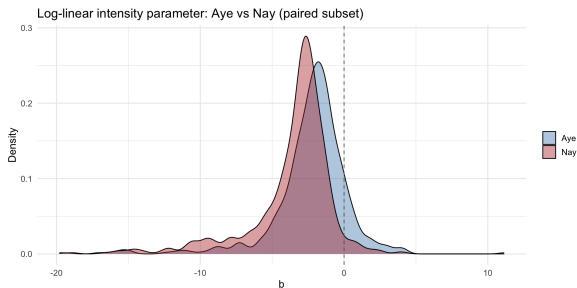


Both rules **relax over time** — the ASM by construction, $\tilde{\phi}$ because residual informativeness vanishes.



But the ASM permits early **approval only**; $\tilde{\phi}$ conditions on $|d_t|$ and can stop in **either** direction.

Is the arrival model tenable?



- Non-homogeneous Poisson, log-linear intensity in normalized time:

$$\lambda(x) = e^{a+bx}, \quad F(x; b) = \frac{e^{bx} - 1}{e^b - 1}$$

- Fits are tight across referenda; participation strongly

front-loaded: pooled median

$$b = -3.05$$

- Rule inputs:

$\Lambda(t) = \Lambda(T) F(t/T; b)$, with
 $\Lambda(T) = 159$, $N = 496$ from the training half

Empirical strategy: strictly out of sample

- Chronological split at the median referendum: **train** (calibrate everything) / **test** (evaluate, frozen parameters)
- Planner states a tolerance α ; the training data pin down the earliest reference date $\bar{t}^*(\alpha)$ whose threshold keeps realized training error within α
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5	21.6	4.5	5.2	19.8	20.0
2	22.9	1.9	1.7	15.4	15.0
1	24.2	0.9	1.3	11.6	11.5

- At $\alpha = 2\%$: median referendum decided on **day 10.8 of 28**, with 60% of final turnout cast (IQR 50–77%), **15 days** faster than the ASM, diverging in 1.7% of test referenda
- $\bar{t}^*$ stable across an order of magnitude in α (days 20.3–24.2): a compact summary of how unrepresentative early participation currently is

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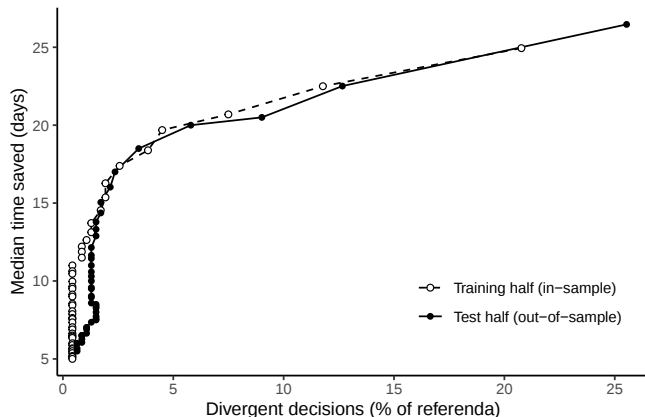
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The speed–accuracy frontier



- Train and test curves nearly coincide
- Steep at the accurate end: the first basis points of error buy nearly all the delay savings; exact zero divergence forfeits the mechanism ($\bar{t} \rightarrow T \Rightarrow$ never stop early)

Takeaways

- Optimal aggregation of asynchronous votes under a finite pool and a hard deadline: a **dynamic majority-quorum rule** with a **lens-shaped** continuation region, collapsing toward the deadline
- Comparative statics governed by the option value of learning
- Large-population limit: one sufficient statistic; a **sequential significance test on informational time**, run by a single error tolerance
- Polkadot, out of sample: **half the decision time at a 2% error tolerance**

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Appendix

Results at a glance

- ① Optimal policy = dynamic majority-quorum rule: unique threshold $d_t^*(s_t)$, collapsing at the deadline (Theorem 1)
- ② Monotone comparative statics in stakes, arrival intensity, and time, driven by the option value of sampling (Proposition 1)
- ③ Large-population, high-stakes limit: closed-form rule on a single statistic; equivalently a deadline-indexed sequential significance test (Theorem 2, Corollary 1)
- ④ Polkadot out of sample: at $\alpha = 2\%$, decision times halved, $< 2\%$ divergence from deadline outcomes

Proof of Theorem 1: architecture

Induction statement: for each s_t , the net stopping gain $G_t(\mu, s) = U(\mu, s) - C_t(\mu, s)$ is non-decreasing in μ on $[\frac{1}{2}, 1]$.

- **Martingale lemma:** $\mathbb{E}[\mu(s^+ + a^+, s^- + a^-) \mid s^+, s^-] = \mu(s^+, s^-)$ for any random batch
(iterated expectations on $\mu = \mathbb{E}[1\{M^+ > N/2\} \mid \cdot]$)
- **FOSD lemma:** for $\mu_{1,t} > \mu_{2,t} \geq \frac{1}{2}$, $P(\mu_{t+1} \leq x \mid \mu_{1,t}, s_t) \leq P(\mu_{t+1} \leq x \mid \mu_{2,t}, s_t)$ for all x
 - likelihood ratio of the predictive for consecutive a^+ : $\frac{k+1+s^+}{k+1} \cdot \frac{a-k}{a-k+s^-}$ increasing in s^+ , decreasing in s^- at fixed $s \Rightarrow \text{MLRP} \Rightarrow \text{FOSD}$, conditional on a
 - averaging over a preserves FOSD because the censored-Poisson batch-size law is independent of μ_t
- **Base case** ($t = T-1$): $G_{T-1} = \kappa - R \mathbb{E}[\max\{0, 1 - 2\mu_T\} \mid \mu_{T-1}, s_{T-1}]$, non-decreasing by FOSD; **inductive step** decomposes G_t into terms inheriting monotonicity

Counterexample: cumulative dominance is not enough (ii)

$N = 5$, $T = 2$, $R = 10$, $\kappa = \frac{1}{5}$; $\lambda = (3, 2)$ dominates $\lambda' = (1, 3)$ cumulatively; state $(s_1^+, s_1^-) = (2, 0)$.

- Continuation at $t = 1$ hinges on the *future* rate only: $\lambda(2) = 2 < 3 = \lambda'(2)$
- Evaluating the thresholds:

$$\lambda(2) = 2 : p_2(3) = 1 - 5e^{-2} \approx 0.323 \leq \frac{2}{5} \Rightarrow \text{stop at } t = 1$$

$$\lambda'(2) = 3 : p_3(3) = 1 - \frac{17}{2}e^{-3} \approx 0.577 > \frac{2}{5} \Rightarrow \text{continue}$$

so $d_1^*(2; \lambda) < d_1^*(2; \lambda')$ despite cumulative dominance of λ

- Volume vs. front-loading: more total arrivals, but fewer *remaining*, lowers the option value now $\Rightarrow$ part (ii) requires dominance over every future sub-interval

Counterexample: back-loading breaks monotone collapse (iii)

$N = 7$, $T = 3$, $\lambda = (1, 1, 5)$, $R = 10$, $\kappa = \frac{1}{2}$; state $(s^+, s^-) = (3, 1)$.

$$U(3, 1) = R \mu(3, 1) = 10 \cdot \frac{13}{14} = \frac{65}{7} \approx 9.286$$

$$C_1(3, 1) \approx 9.197 < U(3, 1) \Rightarrow \text{stop at } t = 1$$

$$C_2(3, 1) \approx 9.411 > U(3, 1) \Rightarrow \text{continue at } t = 2$$

- Hence $d_1^*(4) = 2 < d_2^*(4) = 4$: the threshold *rises* in anticipation of the informative last period
- Front-loading in part (iii) rules exactly this out: each period is then the most informative remaining

Theorem 2, sketch I: the evidence statistic and the information clock

Step 1: Laplace expansion of the posterior. Posterior of M^+ is log-concave BetaBinomial($N - s, s^+ + 1, s^- + 1$); define the evidence statistic as the log-posterior gap between mode $\hat{m}$ and tie state $m_0 = N/2$: $\ell := l(\hat{m}) - l(m_0)$.

- Taylor at the mode ($l'(\hat{m}) = 0$): with $\sigma^2 = \frac{N(N-s)}{4s}(1 + o(1))$ and $m_0 - \hat{m} = -\frac{Nd}{2s} + O(1)$,
 $2\ell = \frac{d^2}{s} \cdot \frac{N}{N-s} (1 + o(1)) =: Z^2$
- Cubic-to-quadratic remainder ratio $\asymp |\gamma|Z \lesssim \log \rho / N \rightarrow 0$ under $\log \rho \leq \varepsilon N^{1/3}$; absolute exponent error $\lesssim (\log \rho)^2 / N \rightarrow 0$
- Summing the Gaussian window: $\mu - \frac{1}{2} = (\Phi(Z) - \frac{1}{2})(1 + o(1)) + O(N^{-1/2})$, $Z = \text{sgn}(d)\sqrt{2\ell}$; log-concavity bounds the tails by $\rho^{-C^2/3}$

Step 2: the information clock. Local alternatives $M^+ = N/2 + \delta$, $\theta := 2\delta/\sqrt{N}$ ($\theta = O(1)$ is the non-trivial range).

- $M_s := (\tilde{d}_s - 2\delta)/(N - s)$ is an **exact martingale** given M^+ : kills the mean-reverting pull-back of sampling without replacement
- Its quadratic variation accumulates at unit rate in $\tau(s) := \frac{s}{N-s}$, the **information clock**; rescaled lead $\hat{X}_\tau := \sqrt{N} \tilde{d}_s / (N - s)$ satisfies $\hat{X}_\tau \mid \theta \sim N(\theta\tau, \tau)$, and $Z_\tau^2 = \hat{X}_\tau^2 / \tau = 2\ell$

Theorem 2, sketch II: belief dynamics and bracketing

Step 3: belief dynamics. Flat prior on $\theta \Rightarrow \theta \mid \mathcal{F}_\tau \sim N(\hat{X}_\tau/\tau, 1/\tau)$: the clock is the posterior **precision**; correctness probability upon stopping is $\Phi(|Z_\tau|)$.

- Remaining learning budget: $\text{Var}(m_{\tau_T} - m_\tau \mid \mathcal{F}_\tau) = \frac{1}{\tau} - \frac{1}{\tau_T}$, known at τ , vanishing at the deadline
- Lead in standard deviations of remaining learning: $\beta := \frac{|m_\tau|}{\sqrt{1/\tau - 1/\tau_T}}$, so the deadline verdict reverses the leader with probability exactly $\Phi(-\beta)$, and $\beta^2 = \frac{(d_t/s_t)^2}{1/s_t - 1/\hat{s}_T}$ — **the statistic of $\hat{\phi}$**

Step 4: bracketing ϕ^* . Deadline option value $\mathcal{G}_N := \mathbb{E}[\max\{0, 1 - 2\mu_{\tau_T}\} \mid \mathcal{F}_\tau]$.

- Doob optional sampling (bounded belief martingale) + Jensen: every stopping time's expected gain over stopping now is $\leq \mathcal{G}_N \Rightarrow$ stop if $R\mathcal{G}_N < \kappa$; continue if $R\mathcal{G}_N > \kappa(T - t)$ — both exact for ϕ^*
- Gaussian evaluation $\mathcal{G}_N = \mathcal{G}(\beta, \Delta\tau/\tau)(1 + o(1))$ uniformly on the relevant window (Poisson averaging over arrivals contributes $1 + o(1)$)
- MLRP of the predictive $\Rightarrow \mathcal{G}_N$ non-increasing in $\beta \Rightarrow$ the stopping set is an up-set: ϕ^* is a threshold rule in β , with cutoff bracketed by the roots $\beta_{T-t}(t) \leq \beta^*(t) \leq \beta_1(t)$ of $\mathcal{G} = c/\rho$, $c \in \{T - t, 1\}$

Theorem 2, sketch III: evaluating the bracket

Step 5: Laplace evaluation at the reversal edge. Re-centering the reversal integral at $\xi = -\beta$:

$$\mathcal{G}(\beta, \frac{\Delta\tau}{\tau}) = \frac{1}{\pi} \sqrt{\frac{\Delta\tau}{\tau}} \frac{e^{-\beta^2/2}}{\beta^2} (1 + o(1))$$

- The prefactor β^{-2} vs. β^{-1} in $\Phi(-\beta)$: the option value is an order β *smaller* than the reversal probability — only marginal reversals pay — the source of the $-2 \log \log \rho$ correction
- Inverting $\mathcal{G} = c/\rho$ by iterated logarithms:

$$\beta_c(t)^2 = 2 \log \rho - 2 \log \log \rho - 2 \log c + \log \frac{\Delta\tau}{\tau} - 2 \log \pi - 2 \log 2 + o(1)$$

- Bracket width: $\beta_1^2 - \beta_{T-t}^2 = 2 \log(T-t) + o(1) = O(1)$ (T fixed); both ends equal $L_\rho + O(1)$ with $L_\rho := 2 \log \rho - 2 \log \log \rho$
- The $O(1)$ band K_ε satisfies $K_\varepsilon \leq L_\rho \frac{\log \log \rho}{\log \rho}$ for large ρ : ϕ^* and $\hat{\phi}$ agree wherever $|\beta^2/L_\rho - 1| > \log \log \rho / \log \rho$ — the band of the theorem
- Posted form: $N/(N - s_t)$ cancels inside β^2 ; carrying it through and evaluating the threshold on the mean path $s_t \rightarrow \Lambda(t)$ yields $\tilde{\phi}$

Data construction and the conviction wedge

- From 1,824 raw referenda: keep substantive dispositions on six focus tracks; drop proposer-instructed rejections (faulty metadata) and degenerate tally paths (too few snapshots, single jumps, large non-monotonicities) → 952
- Exclude Root track from the stopping evaluation (single-slot queue distorts timing) → 933 (467 train / 466 test)
- Delegated votes register in bulk when the delegate votes; arrivals measured by directly voting accounts
- Conviction-weighted ASM outcome vs. final account majority: disagreement in $\approx 17\%$ of referenda; hence errors are benchmarked against the *deadline account majority*, which any account-based rule could at best replicate

Caveat: would behavior change under $\tilde{\phi}$?

- The counterfactual replays historical paths; adopting $\tilde{\phi}$ changes timing incentives (Lucas critique)
- The rule's inference needs only that **arrival timing is uninformative about preferences**: a *symmetric* shift (everyone moves earlier) leaves the tally representative; recalibrate $\Lambda(\cdot)$ and the rule survives
- Structural reason to expect symmetry: $\tilde{\phi}$ conditions on $|d_t| \Rightarrow$ early acceptance *and* early rejection, neutral between alternatives; incentives attach to being ahead or behind, not to the alternative favored
- The incumbent ASM allows early *approval only* $\Rightarrow$ adopting $\tilde{\phi}$ **removes** a built-in asymmetry rather than introducing one
- Still a hypothesis, not a theorem: validation requires deliberately generated variation (randomized stopping rules, staged low-stakes deployment with the deadline as backstop)

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